

# Robust Vetoes for GW Burst Triggers Using Known Instrumental Couplings

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# NOISE TRANSFER

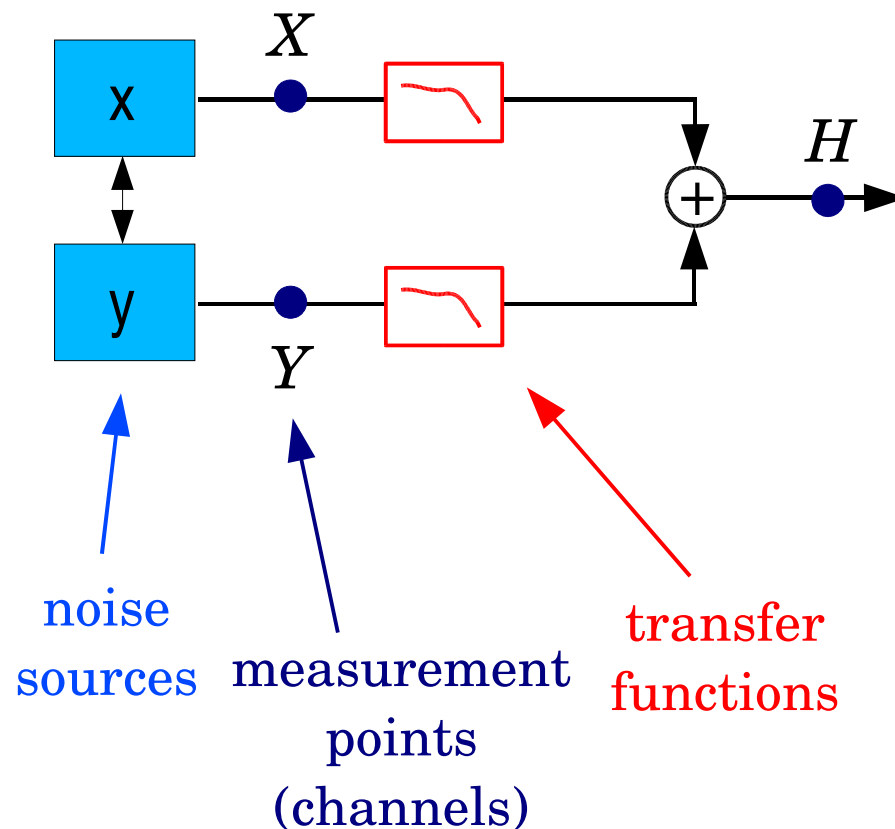


- Transfer function from  $X$  to  $H$  can be measured by injecting noise at  $X$  and measuring  $X(f)$  and  $H(f)$  [1]

$$T_{XH}(f) = \frac{\tilde{X}(f)^* \tilde{H}(f)}{\tilde{X}(f)^* \tilde{X}(f)}$$

- Assuming that the system is linear and time-invariant, the noise in  $X$  can be *transferred (mapped)* to  $H$  by

$$\tilde{X}'(f) = T_{XH}(f) \tilde{X}(f)$$



# VECTOR SPACE PICTURE

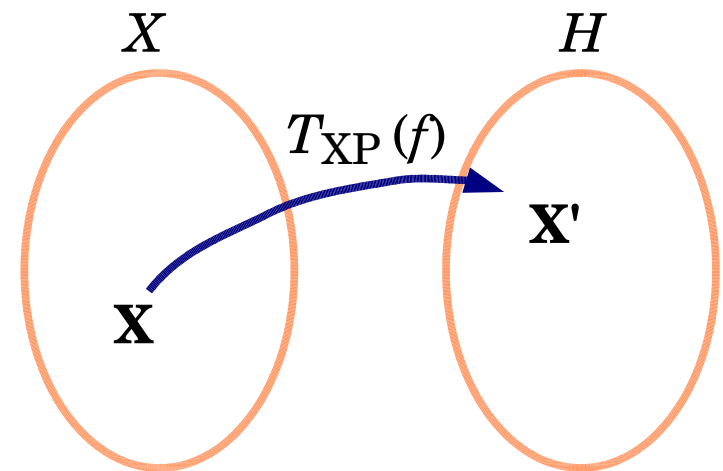


- $X(f)$  and  $H(f)$  can be thought of as components of two vectors  $\mathbf{X}$  and  $\mathbf{H}$  defined in two different Hilbert spaces.

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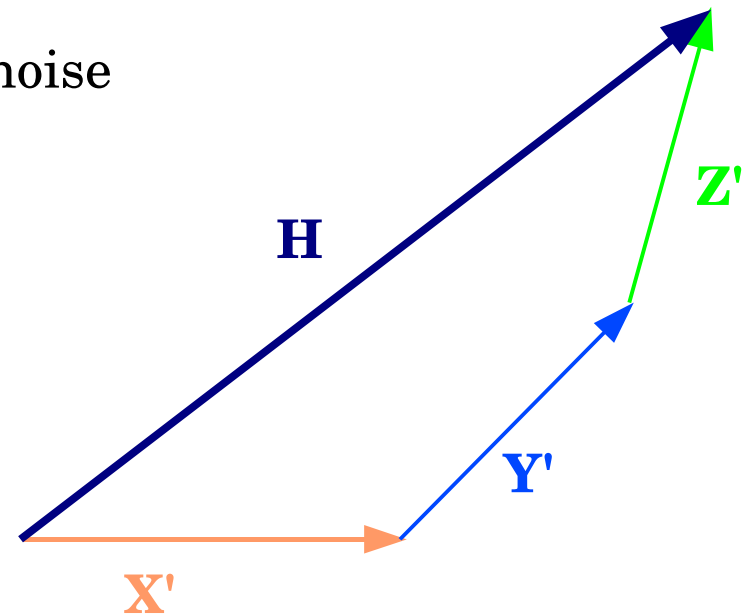
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- $\mathbf{H}$  is made up of many such 'mapped' noise vectors.

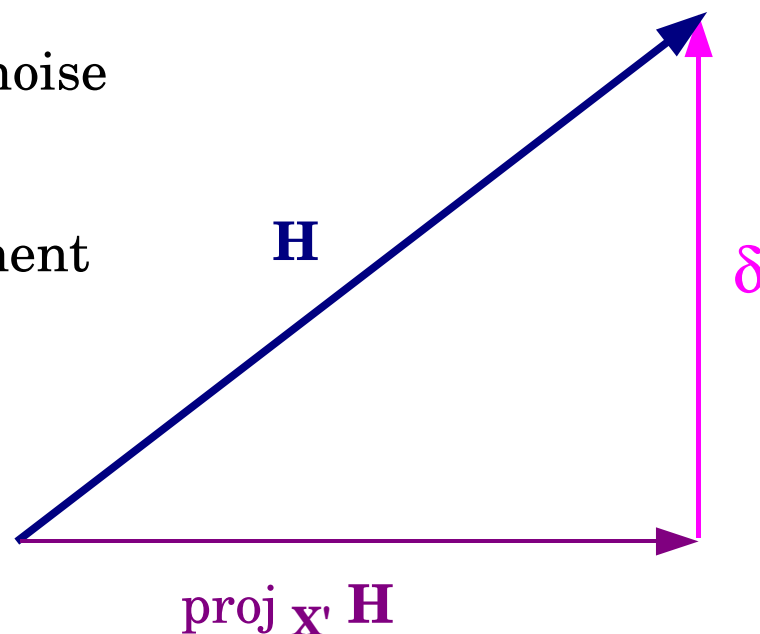


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- $\mathbf{H}$  is made up of many such 'mapped' noise vectors.
- Given  $\mathbf{X}'$ , one can calculate the component of  $\mathbf{H}$  that is orthogonal to  $\mathbf{X}'$  by a Gram-Schmidt orthogonalization

$$\tilde{\delta} = \tilde{\mathbf{H}} - \text{proj}_{\tilde{\mathbf{X}}'} \tilde{\mathbf{H}}$$



# VETO STRATEGY



- If a non-stationarity (glitch) originates from channel  $X$ , it changes the statistics of that segment of data in channel  $X$ , and hence, in channel  $H$ .
- **But, the statistics of the component of  $\mathbf{H}$  that is orthogonal to  $\mathbf{X}'$  remain unchanged.**
- This can be tested by a statistical hypothesis test.
- We take the set of coincident triggers in channels  $X$  and  $H$  and compute  $\delta$  from the segment of the data containing the burst.
- If  $\delta$  at the time of the burst is statistically the same as at other times  $\Rightarrow$  the non-stationarity is originated from channel  $X$ . Thus, we veto the trigger.

# TEST STATISTIC



- Construct the 'excess-power' statistic [2] from  $\delta$

$$\epsilon = \sum_{k=m}^{m+M} P_k, \quad P_k = \frac{|\tilde{\delta}_k|^2}{\sigma_k^2}.$$

$\delta$  in  $k$ th frequency bin

Expected variance of  $\delta$  in  $k$ th frequency bin

[2] W G Anderson et al. PRD **63**, 042003 (2001).



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- Threshold  $\tau$  giving a rejection probability  $\psi$  can be calculated from

$$\psi = \int_0^{\tau} \Gamma(x; \alpha, \beta) dx$$

Prob. density of Gamma  
dist. with scale parameter  $\alpha$   
and shape parameter  $\beta$

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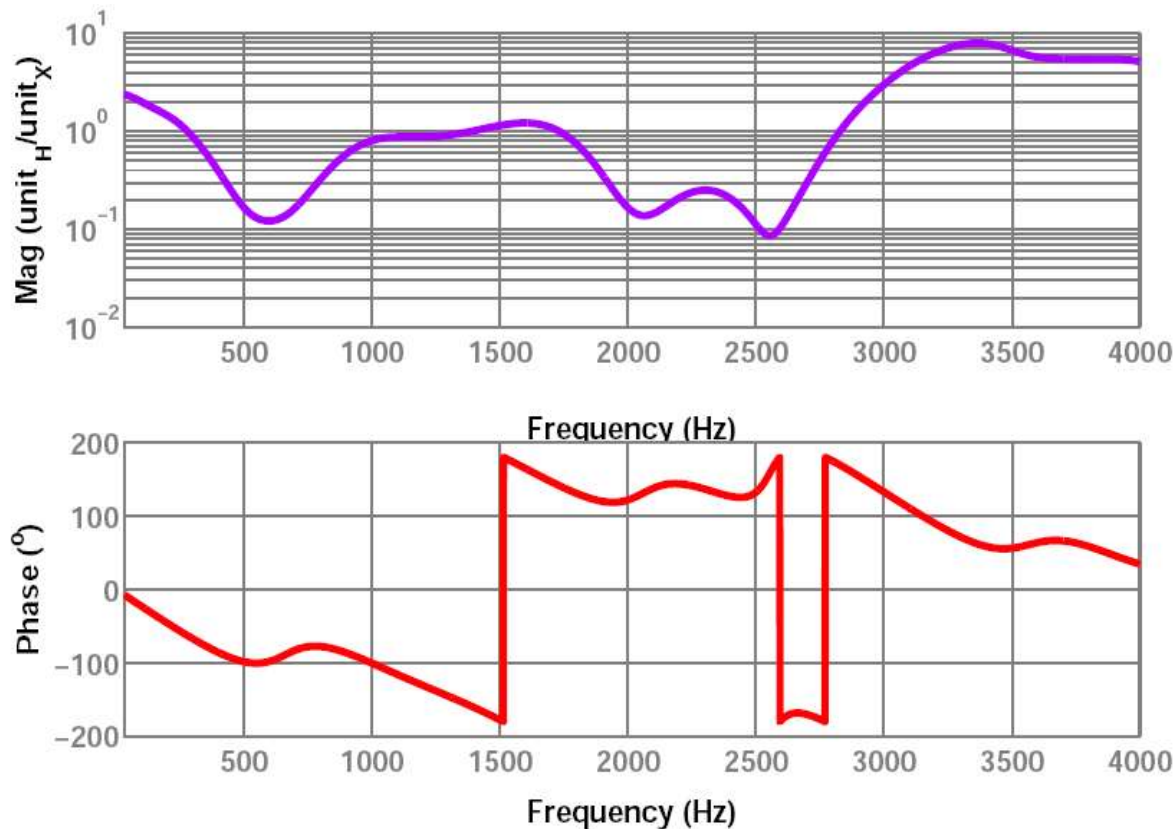
$$\psi = \int_0^{\tau} \Gamma(x; \alpha, \beta) dx$$

- $\alpha$  and  $\beta$  are estimated from stationary data around the glitch.

# SOFTWARE INJECTIONS



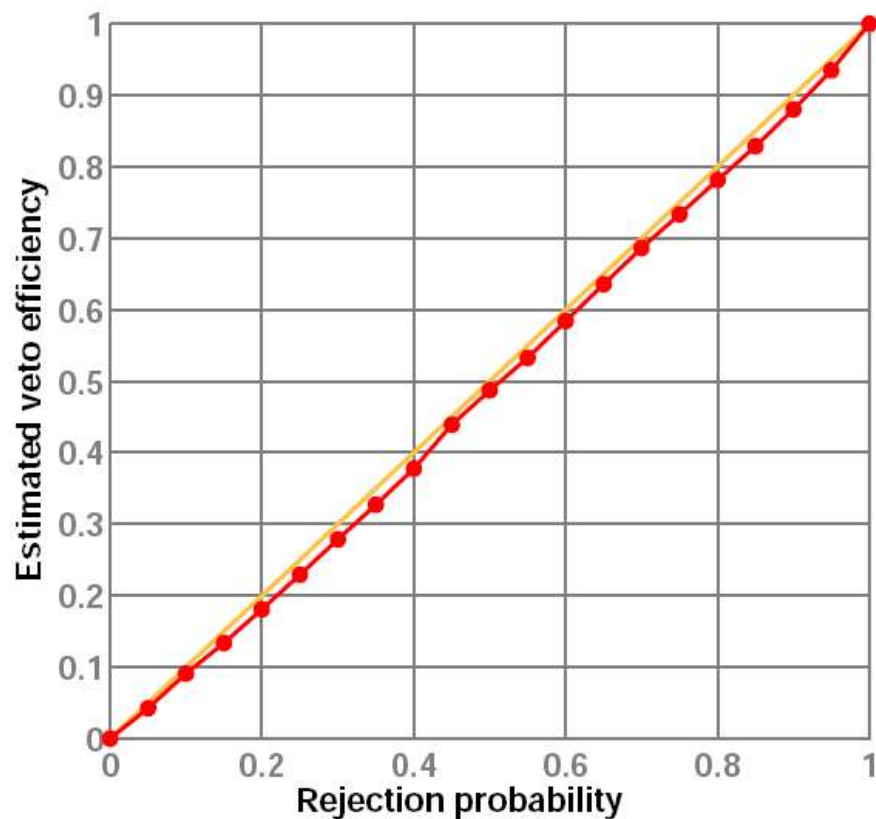
- Q9 sine-Gaussians are injected to white-noise – channel  $X$ .
- To get channel  $H$ : Channel  $X$  data is filtered using a time-domain filter, then some extra noise is added to simulate other components of  $H$ .
- The response of the filter is the transfer function from  $X$  to  $H$ .



# SOFTWARE INJECTIONS



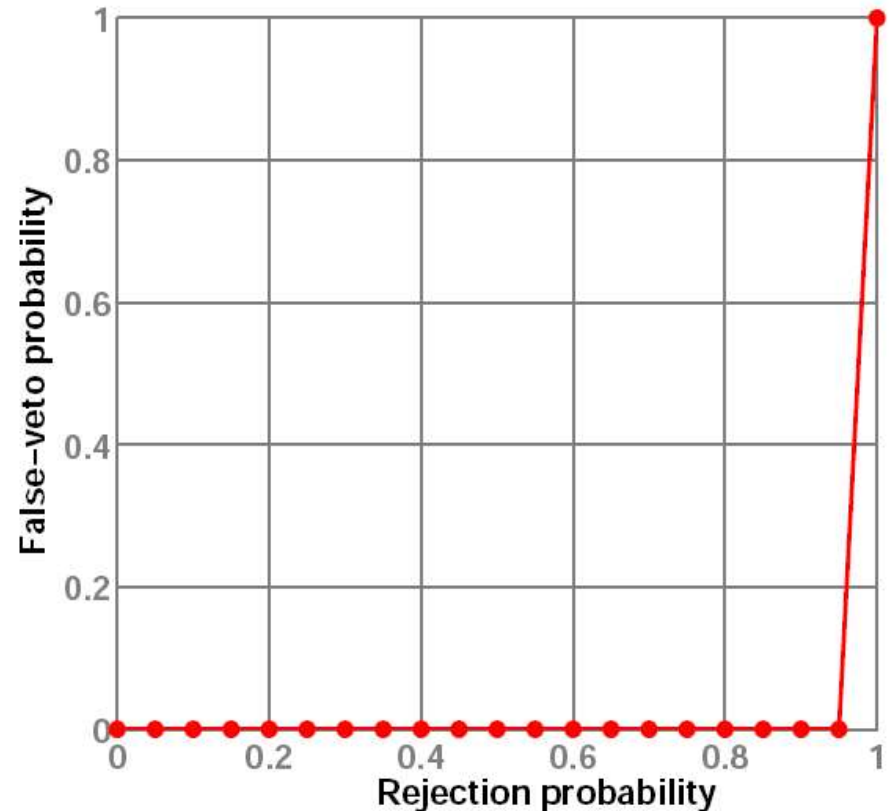
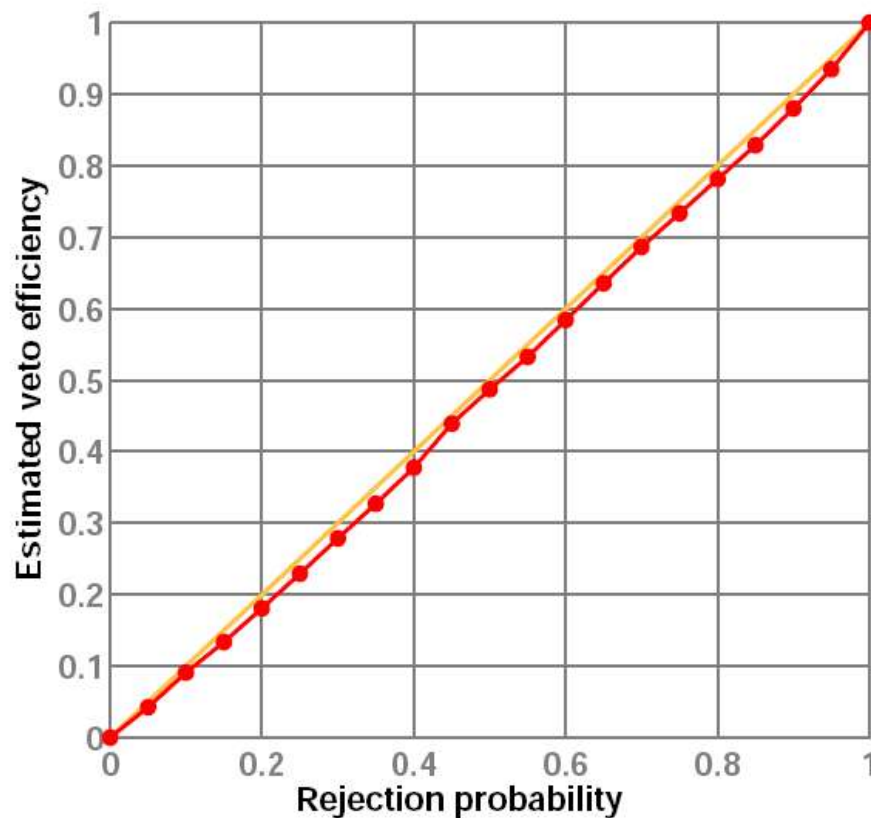
- Perform the veto analysis after choosing different thresholds. Compare the estimated veto fraction with the predicted rejection probability corresponding to each threshold.



# SOFTWARE INJECTIONS



- Perform the veto analysis after choosing different thresholds. Compare the estimated veto fraction with the predicted rejection probability corresponding to each threshold.
- A plausible estimation of false-veto probability: Inject SG waveforms with random parameters to  $X$  and  $H$ . Perform the veto analysis.



# AN ALTERNATIVE METHOD: 'TRIGGER MAPPING'



- An ETG is run over channels  $X$  and  $H$  and two sets of triggers are generated.
- Parameters of the burst triggers in channel  $X$  can be mapped to channel  $H$ , making use of the transfer function from  $X$  to  $H$ .
- If a trigger in  $X$ , mapped to  $H$ , is consistent (in time, frequency and amplitude) with a trigger in channel  $H$ , veto it.
- A less-rigorous, but computationally inexpensive method.



# TRIGGER MAPPING

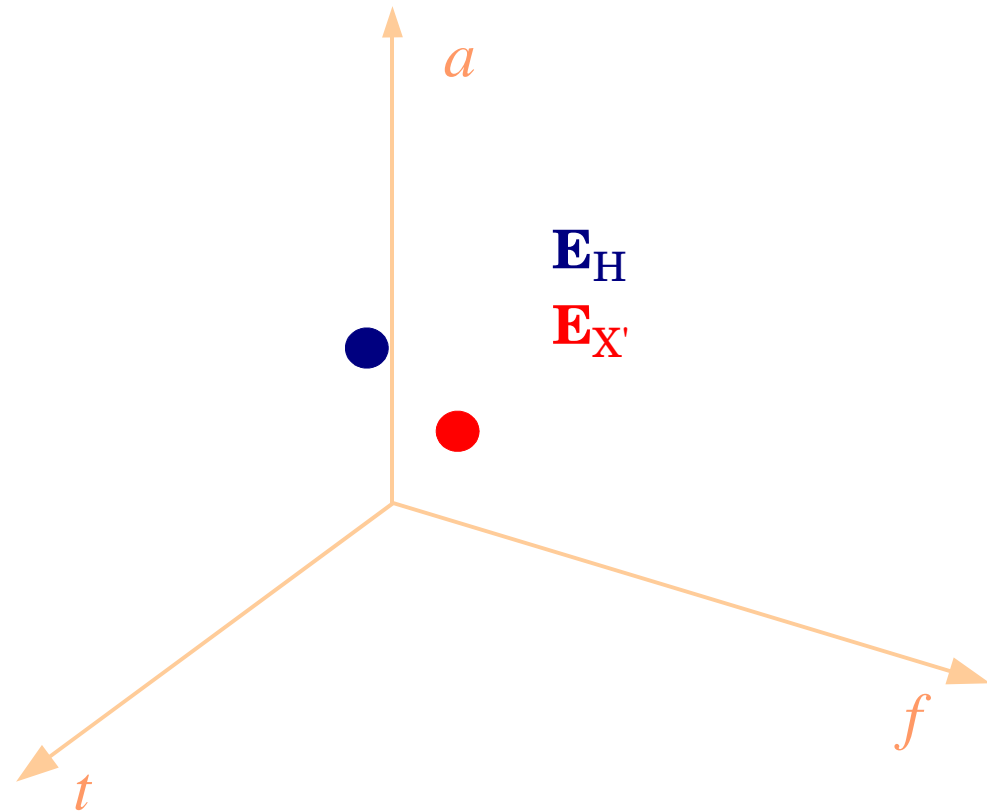


- A certain number of parameters are associated with each event  $\mathbf{E}_X$  in channel  $X$ , like time, central freq, amplitude  $\rightarrow \mathbf{E}_X$  as point  $(E_X^a, E_X^f, E_X^t)$  in a 3-parameter space.
- If we make an assumption about the power spectrum of the underlying burst (like, it is a Gaussian with amplitude  $E_X^a$ , central frequency  $E_X^f$  etc.), we can map the power spectrum to channel  $H$  using the transfer function, and then can re-estimate the parameters from the 'mapped' power spectrum.
- Mapping the point  $\mathbf{E}_X$  (in the space of  $X$ -triggers) to  $\mathbf{E}_X'$  (in the space of  $H$ -triggers).



# IDENTIFYING CONSISTENT EVENTS

If  $\mathbf{E}_X$  is 'sufficiently close' to  $\mathbf{E}_H$ , veto the trigger.



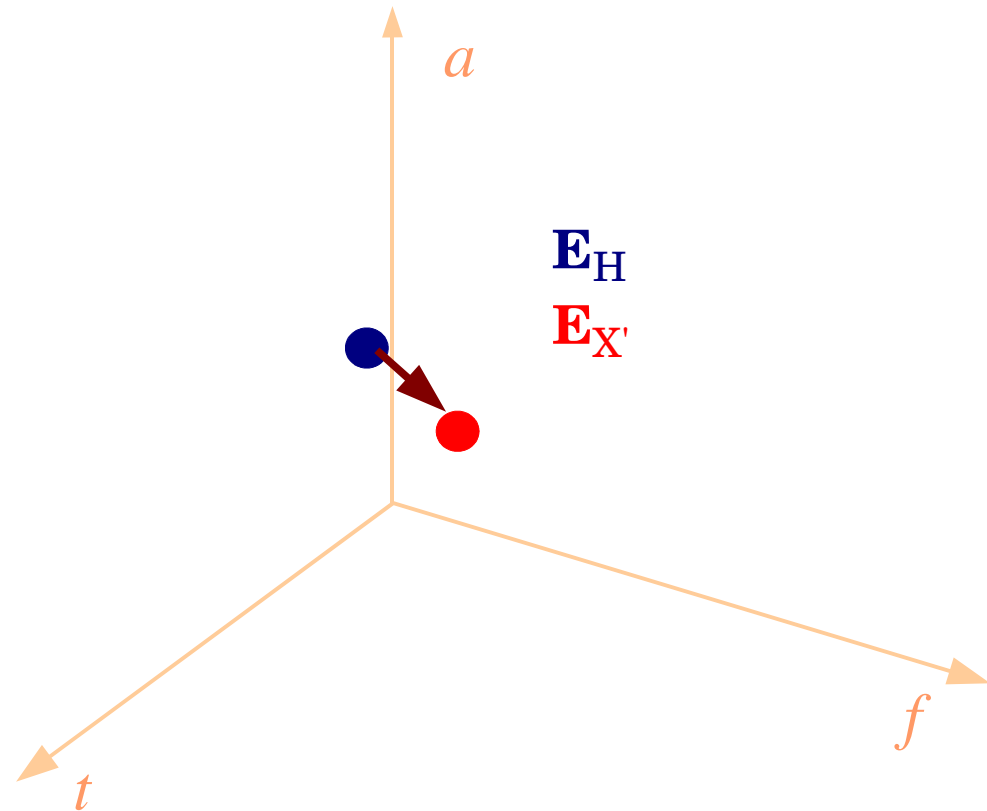


# IDENTIFYING CONSISTENT EVENTS

If  $\mathbf{E}_X$  is 'sufficiently close' to  $\mathbf{E}_H$ , veto the trigger.

- Compute the 'distance vector' between  $\mathbf{E}_X$  and  $\mathbf{E}_H$ :

$$\mathbf{w} \equiv \mathbf{E}_H - \mathbf{E}'_X$$



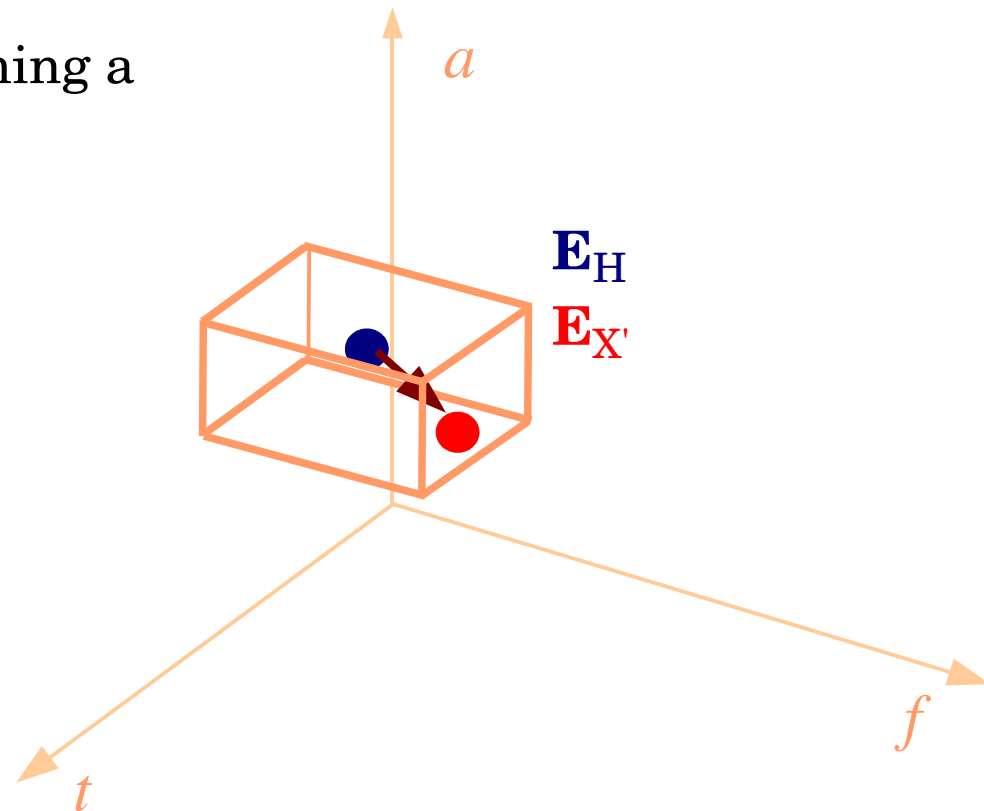
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- If  $\mathbf{w} \leq \tau$ , veto the trigger; thus defining a 'consistency volume'.





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$$\mathbf{w} \equiv \mathbf{E}_H - \mathbf{E}'_X$$

- If  $\mathbf{w} \leq \tau$ , veto the trigger; thus defining a 'consistency volume'.
- If the ETG errors are Gaussian distributed (with zero mean),  $\mathbf{w}$  will be distributed according to Gaussian distributions of mean 0 and variance  $\sigma^2(\mathbf{w})$ .
- Threshold  $\tau$  corresponding to a given rejection probability can be calculated from the prob. densities of Gaussian distributions with mean zero and variance  $\sigma^2(\mathbf{w})$ .

# SOFTWARE INJECTIONS

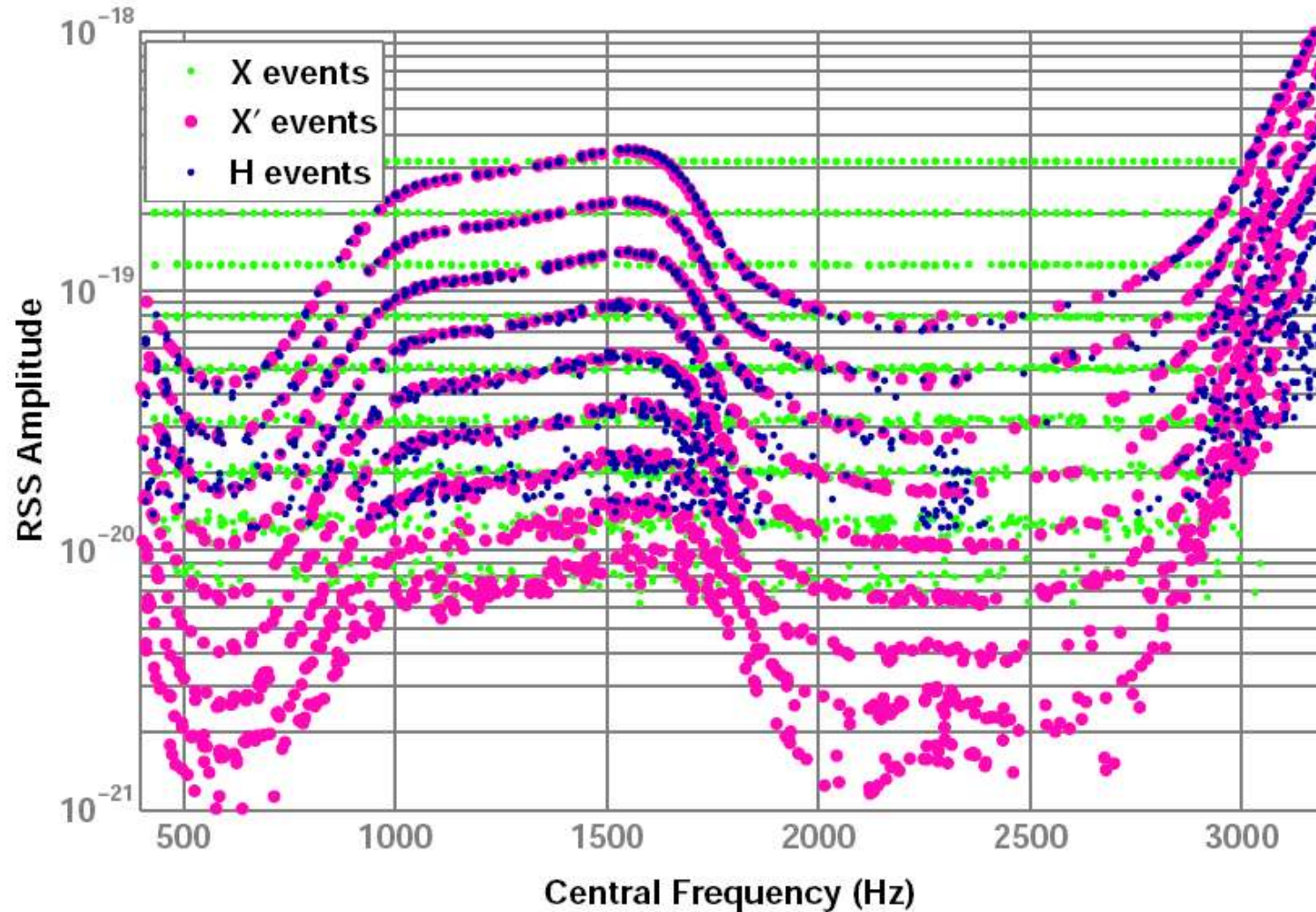


- Data in channels  $X$  and  $H$  are generated as described before.
- The variance of errors in the parameter-estimation of the ETG is estimated by comparing the trigger-parameters with injected parameters.
- Power spectrum of the bursts in channel  $X$  is approximated by a Gaussian. (This is good for SG waveforms. But how good is it in real-life?).
- Veto analysis is performed by choosing different thresholds.

# SOFTWARE INJECTIONS



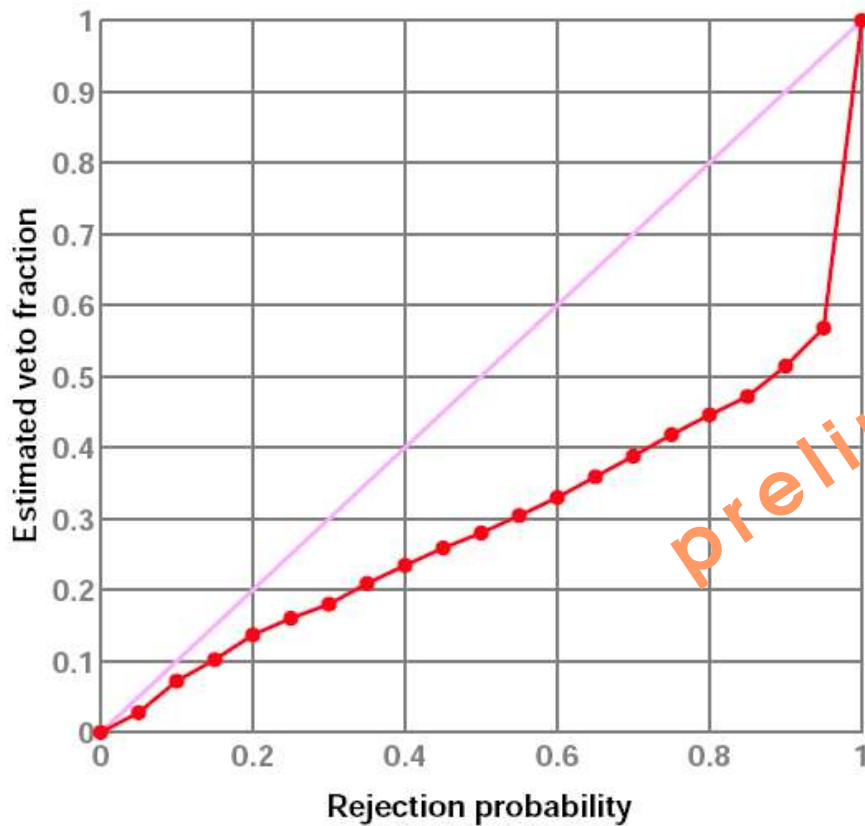
- An example from SG injections:



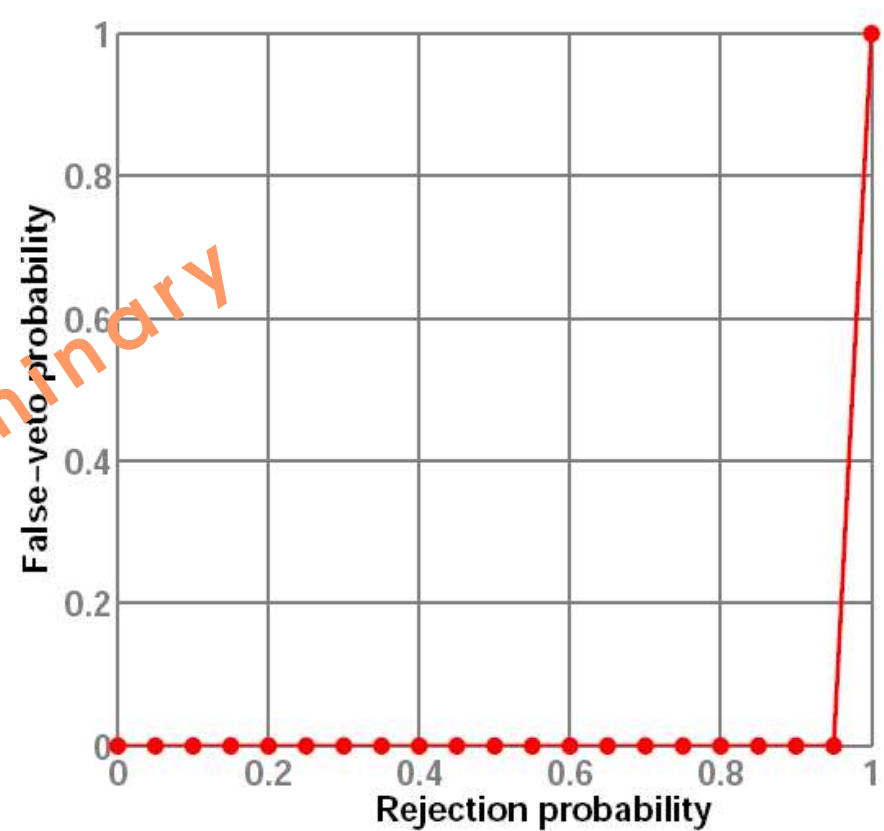
# SOFTWARE INJECTIONS



## Estimating veto efficiency



## Estimating false-veto probability





# S U M M A R Y



- A robust veto strategy is formulated making use of the known instrumental couplings.
- Based on projecting the noise at the detector output into two orthogonal directions, making use of the measured transfer function.
- An alternative method – less rigorous, but computationally inexpensive – is proposed, making use of the trigger parameters estimated by the ETG.
- Work in progress to apply in to GEO600 data.