Searching for Echoes of Gravitational Waves from the Coalescence of Exotic Compact Objects: A Bayesian Approach

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Introduction

Gravitational Waves from Compact Binary Coalescences

Three stages of gravitational waves from CBCs

- Inspiral
- Merger
- Ringdown



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Figure: The numerical relativity simulation (in red) and reconstructed template (in gray) of GW150914. Figure taken from [1].

Echoes from Exotic Compact Objects (ECOs)

- ECOs: Compact objects that aren't White Dwarfs, Neutron Stars nor Black Holes
- Examples of ECOs as alternatives to black holes:
 - Gravastar
 - Fuzzball
- A common feature of these alternatives: Some **structure near the would-be event horizon**, at r = 2M + I, where $I \ll M$

 \Rightarrow Echoes in late-time ringdown phase



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Abedi et al.'s Claim of Evidence

Abedi, Dykaar and Afshordi [2] published a paper on December 2016, claiming that they have found evidence of Planck-scale structure near the black hole event horizons at a combined 2.9σ significance level on GW150914, LVT151012 and GW151226 using matched filtering technique.



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Figure: Best-fit echo template of GW159014 found by Abedi et al. Figure taken from [2].

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Objectives of this project

 The data analysis method adopted by Abedi et al. [2] received a lot of critics, especially the background estimation

▶ In light of the results claimed in [2], we would like to

- 1. Verify their results using their phenomenological model with **Bayesian analysis** instead
- 2. Ultimately, we aim to verify their results using different models (robust evidence), and extract the physics of ECOs

Throughout the project, we will use geometrized unit c = G = 1

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- A model that fits the data best does not imply the model gives the highest evidence
- A more complicated model (i.e. with more free parameters) often fits noisy data better than a simpler model (i.e. with less free parameters)
- This is similar to over-fitting in regression
- Bayesian analysis embodies the Occam's razor and penalizes more complicated models automatically

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And More

- How to sample the posterior probability distribution? Markov Chain Monte Carlo (MCMC)
- How to sample multi-modal distribution well? Affine-invariant ensemble sampler + Parallel tempering
- How to calculate the evidence? Thermodynamic integration
- How to reduce the correlation between samples?
 Burn-in + Thinning by Autocorrelation Length (ACL)
- How to compute the point estimates of the parameters? Maximum Likelihood Estimator (MLE), Maximum A Posteriori Estimator (MAP), Bayes Estimator
- How to estimate the precision of the point estimates? Minimum (Bayesian) Credible Interval
- Refer to my final report for details/Ask me later

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Phenomenological Templates for Echoes

Template by Abedi et al.

Parameter	Description
$\Delta t_{ m echo}$	The time interval between each echo
t _{echo}	The time of arrival of the first echo
t_0	The time of truncation of the GW CBC template
	$\mathcal{M}_{I}(t)$ to produce the echo template $\mathcal{M}_{TE,I}(t)$
γ	The damping factor
A	The (overall) amplitude of the echoes





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Bayesian Analysis on Fake Strain Data

- 1. Generating Fake Strain Data
- 2. Parameter Estimation on Fake Strain Data
- 3. Hypothesis Testing on Fake Strain Data
- 4. Estimation of Statistical Significance of Detection in Colored Gaussian Noise

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Bayesian Analysis on Fake Strain Data

- Before performing the Bayesian analysis on actual strain data, it is educational to perform the same analysis on fake strain data first, namely strain data with colored Gaussian noise and IMR signal with echoes (which we will refer as *IMRE*) of known parameters
- Taken care of sky location and IFO antenna pattern



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Technical details of the injection run

<u>Generation of fake strain</u> IMR approximant Analytical PSD Characteristic SNR in H1 Characteristic SNR in L1

Computational resources used

Run time

Number of CPU cores used Total CPU time

Memory usage

Hyperparameters in MCMC

Number of iterations Number of walkers Number of temperature chains Fraction of burn-in

SEOBNRv4_ROM aLIGOZeroDetHighPower 50.0 38.3 17:18:19 5 3 days 08:44:35 4.58 GB

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Corner Plot

Accurate (close to injected value) and precise (narrow posterior distribution)



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Prior Range and Summary Statistics

Parameter	Prior range		
Α	[0.01, 0.99]		
γ	[0.01, 0.99]		
t ₀ (s)	[-0.05, 0.0]		
t _{echo} (s)	[0, 0.1]		
Δt_{echo} (s)	[0.1, 0.2]		

Table: The prior range of the echo parameters

Parameter	MLE/MAP	Mean	Median	SD	90% CI	Injected Value
Α	0.184	0.180	0.178	0.0272	[0.133, 0.224]	0.23
γ	0.725	0.708	0.706	0.105	[0.548, 0.887]	0.64
t ₀ (s)	-0.0135	-0.0158	-0.0149	0.00427	[-0.0220, -0.0104]	-0.01
t _{echo} (s)	0.0500	0.0499	0.0499	0.000137	[0.0496, 0.0500]	0.05
Δt_{echo} (s)	0.150	0.150	0.150	0.000159	[0.150, 0.150]	0.15

Table: The summary statistics of a parameter estimation run. 90% CI refers to the 90% minimum Bayesian credible interval

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A Closer Examination: Δt_{echo}

- The MAP is very close the injected value (represented) by the vertical dashed line)
- ▶ The 90% Bayesian credible interval ([0.1498, 0.1503]) is much narrower than the prior range ([0.1, 0.2]), that means the range is shrank by about 99.5%.
- The chain converges near the true value and hovers around it.





(e) The estimate of 1D (f) The posterior samples of marginal posterior probability Δt_{echo} density for Δt_{echo} by histogram

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Parameter Estimation on Fake Strain Data

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A Closer Examination: A

- As for amplitude-related parameters A and γ, the parameter estimation is not as accurate and precise as those time-related parameters
- ► The range is not shrank as much as compared with ∆t_{echo}, only by about 90.8%
- The posterior samples for A also did not converge well as compared with Δt_{echo}





(a) The estimate of 1D (b) The posterior samples of marginal posterior probability *A* density for *A* by histogram

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Hypothesis Testing on Fake Strain Data

- Ultimately, we want to perform hypothesis testing on a given strain data, and parameter estimation is just an intermediate step
- ► In the language of Bayesian inference, we have two competing models H₀ and H₁, which are just null hypothesis and alternative hypothesis in frequentist language, and they are

 $H_0: \neg Echo \Rightarrow d = n + h_{IMR}$ $H_1: Echo \Rightarrow d = n + h_{IMRE}$

Note that we assumed that there is a gravitational wave signal in the data, and we are only interested in knowing whether there are echoes in the data or not Searching Echoes

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Hypothesis Testing of the injection run

- The log Bayes factor of IMRE versus IMR is the detection statistic we are going to use to decide whether we claim there is an IMRE signal or IMR signal in the data
- ► Since the log Bayes factor In B^{Echo}_{¬Echo}, and equivalently log odds ratio In O^{Echo}_{¬Echo}¹, is greater than 1, we can conclude, from the Bayesian point of view, that the data favors the existence of echoes, which is indeed the case

	Value
Log Evidence for IMRE in the data, In Z _{IMRE}	-40490.980
Log Evidence for IMR in the data, $\ln Z_{IMR}$	-40529.718
Log Bayes factor for IMRE vs. IMR, In $B_{\neg Echo}^{Echo}$	38.738

Table: A summary of log evidence and log Bayes factor calculated for hypothesis testing

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Estimation of Statistical Significance of Detection in Colored Gaussian Noise

- A natural question to ask is that how significant statistically is the detection statistic
- Simply put, how strong is the evidence² of the detection?
- In the Bayesian school, there are different *empirical scales*, such as Jeffreys' scale, to interpret the strength of the evidence. However, they are **subjective and not universally applicable**
- Therefore, we are not going to use them

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Posterior Probability of Data Containing IMRE Signal

The posterior probability that the data (injected with IMRE signal) contains echoes, given that there is a gravitational wave signal is

 $Pr(Echo | d, GW) = 1 - 1.50 \times 10^{-17}$

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► This tells you that, in the Bayesian framework, the probability of the data containing no echoes is as small as 1.50 × 10⁻¹⁷!

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The Frequentist Approach: p-value

- The Bayesian posterior probability fails to tell us that how likely the so-called evidence is simply due to random background noise, since we only consider merely one set of data!
- ► The frequentist approach can answer the following question: Given the null hypothesis H₀ is true, how likely (i.e. the probability) are the data going to be as *extreme or more extreme than* the observed data?
- The probability that we are looking for is exactly the frequentist p-value
- We need the null distribution of a test statistic T to calculate the p-value

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Sampling the Null Distribution

- To sample the null distribution of T in the case of colored Gaussian noise, we can perform a lot of runs with IMR injection (i.e. assuming that H₀ is true) and different realization of the noise
- In this way, we can see how likely the noise will cause the detection statistic as high as, or higher than, the detected value

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Sampling the Null Distribution

- 8500 such runs were performed
- The detected value in the injection run stands out from the background, *consistent* with the posterior probability approach that the the detected value has a very high statistical significance
- But it also means that it will be difficult to estimate the p-value accurately



Figure: The histogram of the null distribution of T. The vertical dashline corresponds to the detected value ・ロト ・ 日 ト ・ 日 ト

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Gaussian KDE of the Null Distribution



Figure: The estimated probability density function using Gaussian KDE is plotted in blue

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Statistical Significance of the Detected Value

The p-value obtained is

$$p = 9.52 \times 10^{-10}$$

Which means that the detected value in the injection run has a statistical significance of

6.01σ

- Note that the injection is relatively loud and we are using Gaussian noise
- We don't expect such a high significance when we use real LIGO noise instead

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Conclusions

- We successfully recovered the IMRE injection with colored Gaussian noise
- The recovered signal parameters were both accurate and precise
- We can estimate the background and hence the statistical significance of the detection
- This means that
 - 1. The analysis method can **find** an IMRE signal, and **infer** its parameters, if there is one in the real LIGO strain data
 - 2. Most importantly, we can report the statistical significance of a detection
 - 3. The pipeline was properly implemented

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On-going Work and Future Work

On-going:

- $1. \ \mbox{Estimate the background for GW150914}$ event
- 2. Optimize and speed up the pipeline
- Future work:
 - 1. Perform the same analysis on LVT151012, GW151226, GW170104,
 - 2. Repeat the analysis using approximants by Mark et al. and Nakano et al.
 - 3. Implement the approximants in lalsuite such that we can utilize lalinference

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